# Risk Aversion and the Labor Margin in Dynamic Equilibrium Models

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#### Coefficient of Relative Risk Aversion

Suppose a household has preferences:

$$E_0\sum_{t=0}^{\infty}\beta^t u(c_t,I_t),$$

$$u(c_t, l_t) = \frac{c_t^{1-\gamma}}{1-\gamma} - \eta l_t$$

What is the household's coefficient of relative risk aversion?

Intro

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Answer: 0

Intro

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$$u(c_t, l_t) = \frac{c_t^{1-\gamma}}{1-\gamma} - \eta \frac{l_t^{1+\chi}}{1+\chi}$$

What is the household's coefficient of relative risk aversion?

Answer: 
$$\frac{1}{\frac{1}{\gamma} + \frac{1}{\chi}}$$

#### Outline of Presentation

- Define risk aversion rigorously in dynamic equilibrium models
- Derive closed-form expressions
- Show the labor margin can have big effects on risk aversion
- Compute numerical solutions far away from steady state
- Relate risk aversion to asset pricing, stochastic discount factor

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#### See the paper for:

- Epstein-Zin preferences
- internal, external habits

#### A Household

Household preferences:

$$E_t \sum_{\tau=t}^{\infty} \beta^{\tau-t} u(c_{\tau}, l_{\tau}),$$

Flow budget constraint:

$$a_{\tau+1}=(1+r_{\tau})a_{\tau}+w_{\tau}I_{\tau}+d_{\tau}-c_{\tau},$$

No-Ponzi condition:

$$\lim_{T\to\infty}\prod_{\tau=t}^{I}(1+r_{\tau+1})^{-1}a_{T+1}\geq 0,$$

 $\{ \emph{w}_{ au}, \emph{r}_{ au}, \emph{d}_{ au} \}$  are exogenous processes, governed by  $heta_{ au}$ 

#### The Value Function

State variables of the household's problem are  $(a_t; \theta_t)$ .

Let:

$$c_t^* \equiv c^*(a_t; \theta_t),$$

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Value function, Bellman equation:

$$V(a_t;\theta_t) = u(c_t^*,l_t^*) + \beta E_t V(a_{t+1}^*;\theta_{t+1}),$$

where:

$$a_{t+1}^* \equiv (1+r_t)a_t + w_tI_t^* + d_t - c_t^*.$$

#### **Technical Conditions**

**Assumption 1.** The function  $u(c_t, l_t)$  is increasing in its first argument, decreasing in its second, twice-differentiable, and strictly concave.

**Assumption 2.** The value function  $V: X \to \mathbb{R}$  for the household's optimization problem exists and satisfies the Bellman equation

$$V(a_t; \theta_t) = \max_{(c_t, l_t) \in \Gamma(a_t; \theta_t)} u(c_t, l_t) + \beta E_t V(a_{t+1}; \theta_{t+1}).$$

**Assumption 3.** For any  $(a_t; \theta_t) \in X$ , the household's optimal choice  $(c_t^*, l_t^*)$  lies in the interior of  $\Gamma(a_t; \theta_t)$ .

**Assumption 4.** The value function  $V(\cdot; \cdot)$  is twice-differentiable in its first argument. (It then follows that  $c^*$ ,  $l^*$  are differentiable.)

#### Assumptions about the Economic Environment

**Assumption 5.** The household is atomistic.

**Assumption 6.** The household is representative.

**Assumption 7.** The model has a nonstochastic steady state,  $x_t = x_{t+k}$  for k = 1, 2, ..., and  $x \in \{c, l, a, w, r, d, \theta\}$ .

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**Assumption 7'.** The model has a balanced growth path that can be renormalized to a nonstochastic steady state after a suitable change of variables.

Compare:

$$E u(c + \sigma \varepsilon)$$
 vs.  $u(c - \mu)$ 

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$$E u(c + \sigma \varepsilon) \approx u(c) + u'(c)\sigma E[\varepsilon] + \frac{1}{2}u''(c)\sigma^2 E[\varepsilon^2],$$

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$$\mu=\frac{-u''(c)}{u'(c)}\frac{\sigma^2}{2}.$$

Compare:

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Compute:

$$u(c - \mu) \approx u(c) - \mu u'(c),$$
  
 $E u(c + \sigma \varepsilon) \approx u(c) + \frac{1}{2} u''(c) \sigma^2.$ 

$$\mu=\frac{-u''(c)}{u'(c)}\frac{\sigma^2}{2}.$$

Coefficient of absolute risk aversion is defined to be:

$$\lim_{\sigma \to 0} 2\mu(\sigma)/\sigma^2 = \frac{-u''(c)}{u'(c)}.$$

Consider a one-shot gamble in period *t*:

$$a_{t+1} = (1 + r_t)a_t + w_t I_t + d_t - c_t + \sigma \varepsilon_{t+1},$$
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- a<sub>t</sub> (state variable, already known at t)
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Note (\*) is equivalent to gamble over income:

$$a_{t+1} = (1 + r_t)a_t + w_t I_t + (d_t + \sigma \varepsilon_{t+1}) - c_t,$$

or asset returns:

$$a_{t+1} = (1 + r_t + \sigma \tilde{\epsilon}_{t+1}) a_t + w_t l_t + d_t - c_t.$$

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Welfare loss from  $\mu$ :

$$V_1(a_t;\theta_t)\frac{\mu}{(1+r_t)}$$

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Welfare loss from  $\mu$ :

$$\beta E_t V_1(a_{t+1}^*; \theta_{t+1}) \mu$$
.

Loss from  $\sigma$ :

$$\beta E_t V_{11}(a_{t+1}^*; \theta_{t+1}) \frac{\sigma^2}{2}.$$

**Definition 1.** The household's coefficient of absolute risk aversion at  $(a_t; \theta_t)$  is given by  $R^a(a_t; \theta_t) = \lim_{\sigma \to 0} 2\mu(\sigma)/\sigma^2$ .

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**Proposition 1.** The household's coefficient of absolute risk aversion at  $(a_t; \theta_t)$  satisfies

$$R^{a}(a_{t};\theta_{t}) = \frac{-E_{t}V_{11}(a_{t+1}^{*};\theta_{t+1})}{E_{t}V_{1}(a_{t+1}^{*};\theta_{t+1})}.$$

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Evaluated at the nonstochastic steady state, this simplifies to:

$$R^{a}(a;\theta) = \frac{-V_{11}(a;\theta)}{V_{1}(a;\theta)}.$$

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## Solve for $V_1$ and $V_{11}$

Benveniste-Scheinkman:

$$V_1(a_t; \theta_t) = (1 + r_t) u_1(c_t^*, I_t^*). \tag{*}$$

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Benveniste-Scheinkman:

$$V_1(a_t;\theta_t) = (1 + r_t) u_1(c_t^*, I_t^*). \tag{*}$$

Differentiate (\*) to get:

$$V_{11}(a_t; \theta_t) = (1 + r_t) \left[ u_{11}(c_t^*, l_t^*) \frac{\partial c_t^*}{\partial a_t} + u_{12}(c_t^*, l_t^*) \frac{\partial l_t^*}{\partial a_t} \right].$$

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Differentiate to get:

$$\frac{\partial I_t^*}{\partial a_t} = -\frac{\lambda_t}{\partial a_t} \frac{\partial c_t^*}{\partial a_t},$$

$$\frac{\lambda_t}{u_{22}(c_t^*, l_t^*) + u_{12}(c_t^*, l_t^*)}{u_{22}(c_t^*, l_t^*) + w_t u_{12}(c_t^*, l_t^*)}.$$

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Household Euler equation:

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Household Euler equation:

$$u_1(c_t^*, I_t^*) = \beta E_t(1 + r_{t+1}) u_1(c_{t+1}^*, I_{t+1}^*),$$

Differentiate, substitute out for  $\partial l_t^*/\partial a_t$ , and use BC, TVC to get:

$$\frac{\partial c_t^*}{\partial a_t} = \frac{r}{1 + w \lambda}.$$

$$V_1(a; \theta) = (1 + r) u_1(c, l),$$

$$V_1(a;\theta) = (1+r) u_1(c,l),$$

$$V_{11}(a;\theta) = (1+r) \left[ u_{11}(c,l) \frac{\partial c_t^*}{\partial a_t} + u_{12}(c,l) \frac{\partial l_t^*}{\partial a_t} \right],$$

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**Proposition 2.** The household's coefficient of absolute risk aversion in Proposition 1, evaluated at steady state, satisfies:

$$R^{a}(a;\theta) = \frac{-V_{11}(a;\theta)}{V_{1}(a;\theta)} = \frac{-u_{11} + \lambda u_{12}}{u_{1}} \frac{r}{1 + w\lambda}.$$

#### Corollary 3.

$$R^{a}(a;\theta) = \frac{-u_{11} + \lambda u_{12}}{u_{1}} \frac{r}{1 + w\lambda} \leq \frac{-u_{11}}{u_{1}} r.$$

If r < 1, then  $R^a(a; \theta)$  is also less than  $-u_{11}/u_1$ .

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$$u(c_t, I_t) = c_t^{\theta} (\overline{I} - I_t)^{1-\theta}.$$

### Relative Risk Aversion

Consider Arrow-Pratt gamble of general size  $A_t$ :

$$a_{t+1} = (1 + r_t)a_t + w_t I_t + d_t - c_t + A_t \sigma \varepsilon_{t+1},$$
 vs.

$$a_{t+1} = (1 + r_t)a_t + w_t I_t + d_t - c_t - A_t \mu.$$

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$$\frac{-A_t E_t V_{11}(a_{t+1}^*; \theta_{t+1})}{E_t V_1(a_{t+1}^*; \theta_{t+1})}.$$
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Risk aversion coefficient for this gamble:

$$\frac{-A_t E_t V_{11}(a_{t+1}^*; \theta_{t+1})}{E_t V_1(a_{t+1}^*; \theta_{t+1})}.$$
 (\*)

A natural benchmark for  $A_t$  is household wealth at time t.

#### Household Wealth

In DSGE framework, household wealth has more than one component:

- financial assets a<sub>t</sub>
- present value of labor income,  $w_t l_t$
- present value of net transfers, d<sub>t</sub>
- present value of leisure,  $w_t(\bar{l} l_t)$ ?

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Leisure, in particular, can be hard to define, e.g.,

$$u(c_t, l_t) = \frac{c_t^{1-\gamma}}{1-\gamma} - \eta \frac{l_t^{1+\chi}}{1+\chi}$$

and  $\bar{l}$  is arbitrary.

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Different definitions of household wealth lead to different definitions of relative risk aversion.

#### Two Coefficients of Relative Risk Aversion

**Definition 2.** The coefficient of relative risk aversion,  $R^r(a_t; \theta_t)$ , is given by (\*), with  $A_t \equiv (1 + r_t)^{-1} E_t \sum_{\tau=t}^{\infty} m_{t,\tau} (c_{\tau}^* + w_{\tau}(\bar{I} - I_{\tau}^*))$ . In steady state:

$$R^{r}(a;\theta) = \frac{-A V_{11}(a;\theta)}{V_{1}(a;\theta)} = \frac{-u_{11} + \lambda u_{12}}{u_{1}} \frac{c + w(I - I)}{1 + w\lambda}.$$

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**Definition 3.** The consumption-only coefficient of relative risk aversion,  $R^c(a_t; \theta_t)$ , is given by (\*), with  $\tilde{A}_t \equiv (1 + r_t)^{-1} E_t \sum_{\tau=t}^{\infty} m_{t,\tau} c_{\tau}^*$ .

In steady state:

$$R^{c}(a;\theta) = \frac{-\tilde{A} V_{11}(a;\theta)}{V_{1}(a;\theta)} = \frac{-u_{11} + \lambda u_{12}}{u_{1}} \frac{c}{1 + w\lambda}.$$

#### Risk Aversion and the IES

## Corollary 5.

- i)  $R^c(a; \theta)$  and the intertemporal elasticity of substitution are reciprocal if and only if  $\lambda = 0$ ;
- ii)  $R^r(a; \theta)$  and the intertemporal elasticity of substitution are reciprocal if and only if  $\lambda = (\bar{l} l)/c$ .

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- ii)  $R^r(a; \theta)$  and the intertemporal elasticity of substitution are reciprocal if and only if  $\lambda = (\bar{l} l)/c$ .

Proof:

$$IES = \frac{-u_1}{u_{11} - \lambda u_{12}} \frac{1}{c}$$

$$R^c(a; \theta) = \frac{-u_{11} + \lambda u_{12}}{u_1} \frac{c}{1 + w\lambda}$$

$$R^r(a; \theta) = \frac{-u_{11} + \lambda u_{12}}{u_1} \frac{c + w(\bar{l} - l)}{1 + w\lambda}$$

Period utility function:  $u(c_t, I_t) = \frac{\left(c_t^{\chi} (1 - I_t)^{1 - \chi}\right)^{1 - \gamma}}{1 - \gamma}$ 

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Relative risk aversion: 
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$$u_1 = \chi c^{(1-\gamma)\chi-1} (1-I)^{(1-\gamma)(1-\chi)} u_2 = -(1-\chi) c^{(1-\gamma)\chi} (1-I)^{(1-\gamma)(1-\chi)-1}$$

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Relative risk aversion: 
$$R^{r}(a; \theta) = \frac{-u_{11} + \lambda u_{12}}{u_{1}} \frac{c + w(1 - l)}{1 + w\lambda}$$

$$u_1 = \chi c^{(1-\gamma)\chi-1} (1-I)^{(1-\gamma)(1-\chi)} u_2 = -(1-\chi) c^{(1-\gamma)\chi} (1-I)^{(1-\gamma)(1-\chi)-1}$$

$$w = \frac{-u_2}{u_1} = \frac{1-\chi}{\chi} \frac{c}{1-I}$$

Period utility function: 
$$u(c_t, I_t) = \frac{\left(c_t^{\chi} (1 - I_t)^{1 - \chi}\right)^{1 - \gamma}}{1 - \gamma}$$

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$$\begin{array}{ll} u_{11} &= \chi \left[ ((1-\gamma)\chi - 1) \right] c^{(1-\gamma)\chi - 2} (1-I)^{(1-\gamma)(1-\chi)} \\ u_{12} &= -\chi (1-\chi)(1-\gamma) c^{(1-\gamma)\chi - 1} (1-I)^{(1-\gamma)(1-\chi) - 1} \end{array}$$

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 $u_1 = \chi c^{(1-\gamma)\chi-1} (1-I)^{(1-\gamma)(1-\chi)}$ 

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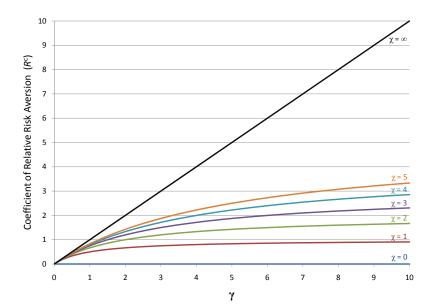
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$$R^c(a;\theta) = \frac{\gamma}{1 + \frac{\gamma}{\gamma} \frac{wl}{c}} \approx \frac{1}{\frac{1}{\gamma} + \frac{1}{\gamma}}, \text{ using } c = wl + ra + d \approx wl.$$

## Example 2



$$u(c_t, l_t) = \frac{c_t^{1-\gamma}}{1-\gamma} - \eta \frac{l_t^{1+\chi}}{1+\chi}, \qquad \gamma = 2, \ \chi = 1.5$$

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Standard RBC model:

$$Y_t = A_t K_t^{1-\alpha} L_t^{\alpha}$$

$$K_{t+1} = (1-\delta)K_t + Y_t - C_t$$

$$C_t^{-\gamma} = \beta E_t (1+r_{t+1})C_{t+1}^{-\gamma}$$

$$\eta L_t^{\chi}/C_t^{-\gamma} = w_t$$

$$r_t = (1-\alpha)Y_t/K_t - \delta$$

$$w_t = \alpha Y_t/L_t$$

$$\log A_t = \rho \log A_{t-1} + \varepsilon_t$$

$$R^{a}(a_{t};\theta_{t}) = \frac{-E_{t}V_{11}(a_{t+1}^{*};\theta_{t+1})}{E_{t}V_{1}(a_{t+1}^{*};\theta_{t+1})}$$

$$R^{a}(a_{t}; \theta_{t}) = \frac{-E_{t}V_{11}(a_{t+1}^{*}; \theta_{t+1})}{E_{t}V_{1}(a_{t+1}^{*}; \theta_{t+1})}$$
$$V_{1}(a_{t}; \theta_{t}) = (1 + r_{t}) u_{1}(c_{t}^{*}, l_{t}^{*})$$

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$$V_{1}(a_{t};\theta_{t}) = (1+r_{t})u_{1}(c_{t}^{*},I_{t}^{*})$$

$$V_{11}(a_{t};\theta_{t}) = (1+r_{t})\left[u_{11}(c_{t}^{*},I_{t}^{*})\frac{\partial c_{t}^{*}}{\partial a_{t}} + u_{12}(c_{t}^{*},I_{t}^{*})\frac{\partial I_{t}^{*}}{\partial a_{t}}\right]$$

$$R^{a}(a_{t}; \theta_{t}) = \frac{-E_{t}V_{11}(a_{t+1}^{*}; \theta_{t+1})}{E_{t}V_{1}(a_{t+1}^{*}; \theta_{t+1})}$$

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$$\frac{\partial l_{t}^{*}}{\partial a_{t}} = -\lambda_{t} \frac{\partial c_{t}^{*}}{\partial a_{t}}$$

$$R^{a}(a_{t};\theta_{t}) = \frac{-E_{t}V_{11}(a_{t+1}^{*};\theta_{t+1})}{E_{t}V_{1}(a_{t+1}^{*};\theta_{t+1})}$$

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$$\frac{\partial I_{t}^{*}}{\partial a_{t}} = -\lambda_{t}\frac{\partial c_{t}^{*}}{\partial a_{t}}$$

$$\frac{\partial c_{t}^{*}}{\partial a_{t}} = \beta E_{t}(1+r_{t+1})\frac{\partial c_{t+1}^{*}}{\partial a_{t}}$$

$$R^{a}(a_{t};\theta_{t}) = \frac{-E_{t}V_{11}(a_{t+1}^{*};\theta_{t+1})}{E_{t}V_{1}(a_{t+1}^{*};\theta_{t+1})}$$

$$V_{1}(a_{t};\theta_{t}) = (1+r_{t})u_{1}(c_{t}^{*},I_{t}^{*})$$

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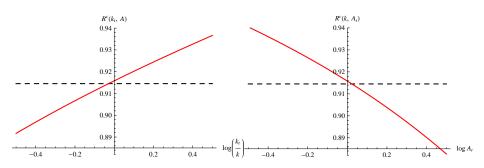
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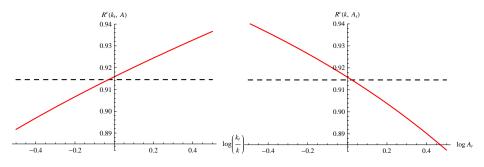
$$\frac{\partial l_{t}^{*}}{\partial a_{t}} = -\lambda_{t}\frac{\partial c_{t}^{*}}{\partial a_{t}}$$

$$\frac{\partial c_{t}^{*}}{\partial a_{t}} = \beta E_{t}(1+r_{t+1})\frac{\partial c_{t+1}^{*}}{\partial a_{t+1}^{*}}\left[(1+r_{t}) + w_{t}\frac{\partial l_{t}^{*}}{\partial a_{t}} - \frac{\partial c_{t}^{*}}{\partial a_{t}}\right]$$

#### Numerical solution to the model:

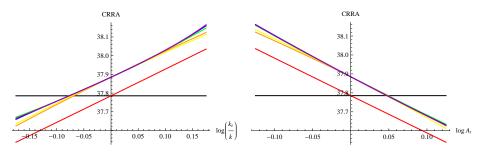


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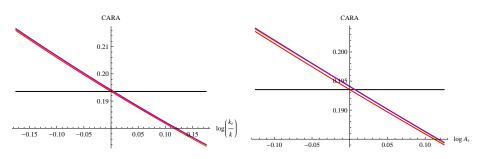


Dashed black line denotes closed-form value of .9143. (Compare to  $\gamma=2$ )

Same exercise with Epstein-Zin Preferences (higher risk aversion):



#### Absolute risk aversion is countercyclical:



Price of an asset at time *t*:

$$E_t m_{t+1} p_{t+1}$$

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Lucas-Breeden stochastic discount factor:

$$m_{t+1} = \frac{\beta u_1(c_{t+1}^*, I_{t+1}^*)}{u_1(c_t^*, I_t^*)}$$

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Recall:

$$E_t m_{t+1} = \frac{1}{1 + r_t^f}$$

Price of an asset at time *t*:

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$$= \frac{-\text{Cov}_t (m_{t+1}, p_{t+1})}{E_t m_{t+1} E_t p_{t+1}}$$

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$$\approx \frac{-\mathsf{Cov}_t (dm_{t+1}, dp_{t+1})}{\beta}$$

$$m_{t+1} = \frac{\beta u_1(c_{t+1}^*, I_{t+1}^*)}{u_1(c_t^*, I_t^*)}$$

$$dm_{t+1} = \frac{\beta}{u_1(c_t^*, l_t^*)} \left[ u_{11}(c_{t+1}^*, l_{t+1}^*) dc_{t+1}^* + u_{12}(c_{t+1}^*, l_{t+1}^*) dl_{t+1}^* \right].$$

$$dm_{t+1} = \frac{\beta}{u_1(c_t^*, l_t^*)} \left[ u_{11}(c_{t+1}^*, l_{t+1}^*) dc_{t+1}^* + u_{12}(c_{t+1}^*, l_{t+1}^*) dl_{t+1}^* \right].$$

Intuitively:

$$dl_{t+1}^* = -\lambda dc_{t+1}^*,$$
  $dc_{t+1}^* = rac{r}{1+w\lambda} da_{t+1},$ 

$$dm_{t+1} = \frac{\beta}{u_1(c_t^*, l_t^*)} \left[ u_{11}(c_{t+1}^*, l_{t+1}^*) dc_{t+1}^* + u_{12}(c_{t+1}^*, l_{t+1}^*) dl_{t+1}^* \right].$$

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$$\frac{-\mathsf{Cov}_t(dm_{t+1},dp_{t+1})}{\beta} = \mathbf{R}^{\mathbf{a}}(\mathbf{a};\theta) \cdot \mathsf{Cov}_t(da_{t+1},dp_{t+1}).$$

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$$+ \frac{-u_1}{u_{11} - \lambda u_{12}} E_{t+1} \sum_{k=1}^{\infty} \frac{1}{(1+r)^k} \left( \frac{r\lambda}{1+w\lambda} dw_{t+k} - \beta dr_{t+k+1} \right).$$

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$$dI_{t+1}^* = -\lambda dC_{t+1}^* + \frac{u_1}{-u_{22} - wu_{12}} dw_{t+1},$$

$$dc_{t+1}^* = \frac{r}{1 + w\lambda} \left[ da_{t+1} + E_{t+1} \sum_{k=1}^{\infty} \frac{1}{(1+r)^k} \left( Idw_{t+k} + dd_{t+k} + adr_{t+k} \right) \right] + \frac{u_1 u_{12}}{u_{11} u_{22} - u_{12}^2} dw_{t+1}$$

$$+ \frac{-u_1}{u_{11} - \lambda u_{12}} E_{t+1} \sum_{k=1}^{\infty} \frac{1}{(1+r)^k} \left( \frac{r\lambda}{1+w\lambda} dw_{t+k} - \beta dr_{t+k+1} \right).$$

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$$dm_{t+1} = \frac{\beta}{u_1(c_t^*, l_t^*)} \left[ u_{11}(c_{t+1}^*, l_{t+1}^*) dc_{t+1}^* + u_{12}(c_{t+1}^*, l_{t+1}^*) dl_{t+1}^* \right].$$

$$dI_{t+1}^* = -\lambda dc_{t+1}^* + \frac{u_1}{-u_{22} - wu_{12}} dw_{t+1},$$

$$dc_{t+1}^* = \frac{r}{1+w\lambda} \left[ da_{t+1} + E_{t+1} \sum_{k=1}^{\infty} \frac{1}{(1+r)^k} \left( Idw_{t+k} + dd_{t+k} + adr_{t+k} \right) \right] + \frac{u_1 u_{12}}{u_{11} u_{22} - u_{12}^2} dw_{t+1}$$

$$+ \frac{-u_1}{u_{11} - \lambda u_{12}} E_{t+1} \sum_{k=1}^{\infty} \frac{1}{(1+r)^k} \left( \frac{r\lambda}{1+w\lambda} dw_{t+k} - \beta dr_{t+k+1} \right).$$

$$dm_{t+1} = \beta \frac{u_{11} - \lambda u_{12}}{u_1} \frac{r}{1 + w\lambda} \left[ da_{t+1} + E_{t+1} \sum_{k=1}^{\infty} \frac{1}{(1+r)^k} \left( ldw_{t+k} + dd_{t+k} + adr_{t+k} \right) \right]$$

$$+\beta E_{t+1} \sum_{l=1}^{\infty} \frac{1}{(1+r)^k} \left(\beta dr_{t+k+1} - \frac{r\lambda}{1+w\lambda} dw_{t+k}\right).$$

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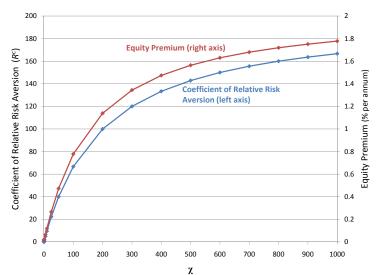
$$R^{r}(a;\theta)\cdot \text{Cov}_{t}(dp_{t+1},\frac{d\hat{A}_{t+1}}{A}) + \text{Cov}_{t}(dp_{t+1},d\Psi_{t+1}),$$

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## **Asset Pricing**

#### Numerical results for an equity claim to consumption, $\gamma =$ 200:



Hansen-Rogerson linear-labor preferences are common:

- Extensive labor margin: Hansen (1985), Rogerson (1988)
- Monetary search: Lagos-Wright (2005)
- Investment: Khan-Thomas (2008), Bachmann-Caballero-Engel (2010), Bachmann-Bayer (2009)

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The present paper suggests ways to model risk neutrality that do not require linear utility of consumption.

## **Empirical Estimates of Risk Aversion**

#### Barsky-Juster-Kimball-Shaprio (1997):

"Suppose that you are the only income earner in the family, and you have a good job guaranteed to give you your current (family) income every year for life. You are given the opportunity to take a new and equally good job, with a 50–50 chance it will double your (family) income and a 50–50 chance that it will cut your (family) income by a third. Would you take the new job?"

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Empirical estimates of risk aversion using methods like these remain valid in the framework of the present paper.

What is different is how these estimates are mapped into model parameters (i.e., risk aversion  $\neq -cu_{11}/u_1$ )

Campbell (1996, 1999): 
$$u(C_t) = \frac{C_t^{1-\gamma}}{1-\gamma}, \quad m_{t+1} = \log \beta - \gamma \Delta c_{t+1}$$

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Country	$E_t(r_{e,t}-r_{f,t})$	$std(\mathit{r}_{e,t}-\mathit{r}_{f,t})$	$std(\Delta c)$	$\gamma$
USA	5.82	17.0	0.91	37.3
JPN	6.83	21.6	2.35	13.4
GER	6.77	20.4	2.50	13.3
FRA	7.12	22.8	2.13	14.6
UK	8.31	21.6	2.59	14.9
ITA	2.17	27.3	1.68	4.7
CAN	3.04	16.7	2.03	9.0

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If 
$$u(C_t) = \frac{C_t^{1-\gamma}}{1-\gamma} - \eta \frac{L_t^{1+\gamma}}{1+\gamma}$$
, then  $\gamma \neq$  risk aversion.

#### Conclusions

- The labor margin has dramatic effects on risk aversion
- Risk aversion is the right concept for asset pricing
- 3 Arrow-Pratt risk neutrality holds for any u with  $u_{11}u_{22} u_{12}^2 = 0$
- Risk aversion and the intertemporal elasticity of substitution are nonreciprocal when there is labor in the model
- Simple, closed-form expressions for risk aversion in DSGE models with:
  - expected utility preferences
  - Epstein-Zin preferences
  - external or internal habits
  - valid away from steady state