Macroeconomic Implications of Changes in the Term Premium

Glenn D. Rudebusch¹ Brian P. Sack² Eric T. Swanson¹

¹Economic Research Federal Reserve Bank of San Francisco

²Macroeconomic Advisers

Frontiers in Monetary Policy Research Federal Reserve Bank of St. Louis October 19, 2006

Outline

- Background and Motivation
- Structural Analysis
- Macro-Finance Analysis
- Reduced-Form Analysis
- Conclusions



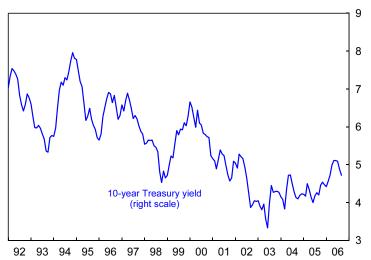
Long-term interest rates have trended lower in recent months even as the Federal Reserve has raised the level of the target federal funds rate by 150 basis points... For the moment, the broadly unanticipated behavior of world bond markets remains a conundrum.

Alan Greenspan, February 2005

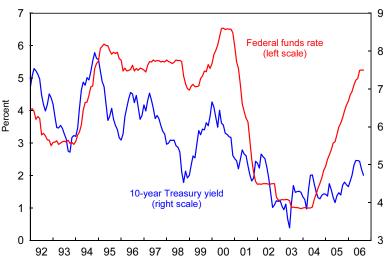
Long-term interest rates have trended lower in recent months even as the Federal Reserve has raised the level of the target federal funds rate by 150 basis points... For the moment, the broadly unanticipated behavior of world bond markets remains a conundrum.

Alan Greenspan, February 2005

Yield on 10-Year US Treasury Securities and Federal Funds Rate



Yield on 10-Year US Treasury Securities and Federal Funds Rate





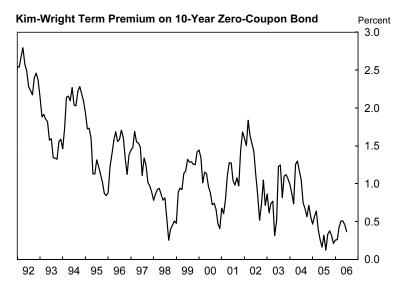


A significant portion of the sharp decline in the ten-year forward one-year rate over the past year appears to have resulted from a fall in term premiums.

Alan Greenspan, July 2005

A significant portion of the sharp decline in the ten-year forward one-year rate over the past year appears to have resulted from a fall in term premiums.

Alan Greenspan, July 2005



Two Questions



Two Questions

What are the macroeconomic implications of a change in the term premium?

Two Questions

What are the macroeconomic implications of a change in the term premium?

How should monetary policy respond to a change in the term premium?



The decline in term premiums in the Treasury market of late may have contributed to keeping long-term interest rates relatively low and, consequently, may have supported the housing sector and consumer spending more generally.

Donald Kohn, July 2005

The decline in term premiums in the Treasury market of late may have contributed to keeping long-term interest rates relatively low and, consequently, may have supported the housing sector and consumer spending more generally.

Donald Kohn, July 2005

The "news" over recent months may instead be the [82 bp] run-up in the ten-year yield [over the past 3 months]... In effect, the FOMC has achieved more tightening of financial conditions over the past three months than it had on net over the entire tightening cycle.

Macroeconomic Advisers, April 2006

The "news" over recent months may instead be the [82 bp] run-up in the ten-year yield [over the past 3 months]... In effect, the FOMC has achieved more tightening of financial conditions over the past three months than it had on net over the entire tightening cycle.

Macroeconomic Advisers, April 2006

To the extent that the decline in forward rates can be traced to a decline in the term premium.... the effect is financially stimulative and argues for greater monetary policy restraint, all else being equal. Specifically, if spending depends on long-term interest rates, special factors that lower the spread between short-term and long-term rates will stimulate aggregate demand. Thus, when the term premium declines, a higher short-term rate is required to obtain the long-term rate and the overall mix of financial conditions consistent with maximum sustainable employment and stable prices.



To the extent that the decline in forward rates can be traced to a decline in the term premium.... the effect is financially stimulative and argues for greater monetary policy restraint, all else being equal. Specifically, if spending depends on long-term interest rates, special factors that lower the spread between short-term and long-term rates will stimulate aggregate demand. Thus, when the term premium declines, a higher short-term rate is required to obtain the long-term rate and the overall mix of financial conditions consistent with maximum sustainable employment and stable prices.



To the extent that the decline in forward rates can be traced to a decline in the term premium.... the effect is financially stimulative and argues for greater monetary policy restraint, all else being equal. Specifically, if spending depends on long-term interest rates, special factors that lower the spread between short-term and long-term rates will stimulate aggregate demand. Thus, when the term premium declines, a higher short-term rate is required to obtain the long-term rate and the overall mix of financial conditions consistent with maximum sustainable employment and stable prices.



To the extent that the decline in forward rates can be traced to a decline in the term premium.... the effect is financially stimulative and argues for greater monetary policy restraint, all else being equal. Specifically, if spending depends on long-term interest rates, special factors that lower the spread between short-term and long-term rates will stimulate aggregate demand. Thus, when the term premium declines, a higher short-term rate is required to obtain the long-term rate and the overall mix of financial conditions consistent with maximum sustainable employment and stable prices.



New Keynesian IS curve (linearized):

$$y_t = \beta E_t y_{t+1} - \frac{1}{\gamma} (i_t - E_t \pi_{t+1}) + \varepsilon_t$$

Conclusions

New Keynesian IS curve (linearized):

$$y_t = \beta E_t y_{t+1} - \frac{1}{\gamma} (i_t - E_t \pi_{t+1}) + \varepsilon_t$$

Solving forward:

$$y_t = -\frac{1}{\gamma} E_t \sum_{j=0}^{\infty} \beta^j (i_{t+j} - \pi_{t+1+j}) + \varepsilon_t$$

New Keynesian IS curve (linearized):

$$y_t = \beta E_t y_{t+1} - \frac{1}{\gamma} (i_t - E_t \pi_{t+1}) + \varepsilon_t$$

Solving forward:

$$y_t = -\frac{1}{\gamma} E_t \sum_{j=0}^{\infty} \beta^j (i_{t+j} - \pi_{t+1+j}) + \varepsilon_t$$

New Keynesian IS curve (linearized):

$$y_t = \beta E_t y_{t+1} - \frac{1}{\gamma} (i_t - E_t \pi_{t+1}) + \varepsilon_t$$

Solving forward:

$$y_t = -\frac{1}{\gamma} E_t \sum_{j=0}^{\infty} \beta^j (i_{t+j} - \pi_{t+1+j}) + \varepsilon_t$$

Note: no role for the term premium in this model

New Keynesian IS curve (linearized):

$$y_t = \beta E_t y_{t+1} - \frac{1}{\gamma} (i_t - E_t \pi_{t+1}) + \varepsilon_t$$

Solving forward:

$$y_t = -\frac{1}{\gamma} E_t \sum_{j=0}^{\infty} \beta^j (i_{t+j} - \pi_{t+1+j}) + \varepsilon_t$$

Note: no role for the term premium in this model Instead, practitioners' model may be more informal:

New Keynesian IS curve (linearized):

$$y_t = \beta E_t y_{t+1} - \frac{1}{\gamma} (i_t - E_t \pi_{t+1}) + \varepsilon_t$$

Solving forward:

$$y_t = -\frac{1}{\gamma} E_t \sum_{j=0}^{\infty} \beta^j (i_{t+j} - \pi_{t+1+j}) + \varepsilon_t$$

Note: no role for the term premium in this model Instead, practitioners' model may be more informal:

IS-LM intuition

New Keynesian IS curve (linearized):

$$y_t = \beta E_t y_{t+1} - \frac{1}{\gamma} (i_t - E_t \pi_{t+1}) + \varepsilon_t$$

Solving forward:

$$y_t = -\frac{1}{\gamma} E_t \sum_{j=0}^{\infty} \beta^j (i_{t+j} - \pi_{t+1+j}) + \varepsilon_t$$

Note: no role for the term premium in this model Instead, practitioners' model may be more informal:

- IS-LM intuition
- Partial equilibrium analysis

In general equilibrium, implications of change in term premium are not clear:

• Why did the term premium change?

- Why did the term premium change?
- Different structural shocks might have different implications for the economy

- Why did the term premium change?
- Different structural shocks might have different implications for the economy
- Term premium might be partly a "wedge"

- Why did the term premium change?
- Different structural shocks might have different implications for the economy
- Term premium might be partly a "wedge"
- Term premium might be related to potential output rather than output gap



- Structural Analysis
 - completely answers question of interest, in principle
 - but number of practical limitations
 - draw some general insights

- Structural Analysis
 - completely answers question of interest, in principle
 - but number of practical limitations
 - draw some general insights
- Macro-Finance Analysis
 - less structural, more tractable
 - more successful empirically
 - but does not address question of interest

- Structural Analysis
 - completely answers question of interest, in principle
 - but number of practical limitations
 - draw some general insights
- Macro-Finance Analysis
 - less structural, more tractable
 - more successful empirically
 - but does not address question of interest
- Reduced-Form Analysis
 - literature using yield curve spread to forecast GDP
 - compare popular term premium measures
 - study importance of term premium for forecasting GDP

Structural Analysis

- Structural Analysis
 - Review Asset Pricing
 - Define Benchmark New Keynesian Model
 - Graph Impulse Responses
 - Discuss Limitations of the Structural Framework



Asset Pricing

Asset pricing:

$$p_t = E_t[m_{t+1}p_{t+1}]$$

Asset pricing:

$$p_t = E_t[m_{t+1}p_{t+1}]$$

Zero-coupon bond pricing:

$$p_t^{(n)} = E_t[m_{t+1}p_{t+1}^{(n-1)}]$$

Asset Pricing

Asset pricing:

$$p_t = E_t[m_{t+1}p_{t+1}]$$

Zero-coupon bond pricing:

$$p_t^{(n)} = E_t[m_{t+1}p_{t+1}^{(n-1)}]$$

$$i_t^{(n)} = -\frac{1}{n} \log p_t^{(n)}$$

Asset pricing:

$$p_t = E_t[m_{t+1}p_{t+1}]$$

Zero-coupon bond pricing:

$$p_t^{(n)} = E_t[m_{t+1}p_{t+1}^{(n-1)}]$$

$$i_t^{(n)} = -\frac{1}{n} \log p_t^{(n)}$$

Notation: let $i_t \equiv i_t^{(1)}$

Background/Motivation

Representative household with preferences:

$$\max E_t \sum_{t=0}^{\infty} \beta^t \left(\frac{(c_t - bC_{t-1})^{1-\gamma}}{1-\gamma} - \chi_0 \frac{I_t^{1+\chi}}{1+\chi} \right)$$

Representative household with preferences:

$$\max E_t \sum_{t=0}^{\infty} \beta^t \left(\frac{(c_t - bC_{t-1})^{1-\gamma}}{1-\gamma} - \chi_0 \frac{I_t^{1+\chi}}{1+\chi} \right)$$

Stochastic discount factor:

$$m_{t+1} = \frac{\beta(C_{t+1} - bC_t)^{-\gamma}}{(C_t - bC_{t-1})^{-\gamma}} \frac{P_t}{P_{t+1}}$$

Representative household with preferences:

$$\max E_t \sum_{t=0}^{\infty} \beta^t \left(\frac{(c_t - bC_{t-1})^{1-\gamma}}{1-\gamma} - \chi_0 \frac{I_t^{1+\chi}}{1+\chi} \right)$$

Macro-Finance Analysis

Stochastic discount factor:

Background/Motivation

$$m_{t+1} = \frac{\beta(C_{t+1} - bC_t)^{-\gamma}}{(C_t - bC_{t-1})^{-\gamma}} \frac{P_t}{P_{t+1}}$$

Parameters: $\beta = .99, b = .66, \gamma = 2, \chi = 1.5$

Continuum of differentiated firms:

• face Dixit-Stiglitz demand with elasticity $\frac{1+\theta}{\theta}$, markup θ

- face Dixit-Stiglitz demand with elasticity $\frac{1+\theta}{\theta}$, markup θ
- set prices in Calvo contracts with avg. duration 4 quarters

- face Dixit-Stiglitz demand with elasticity $\frac{1+\theta}{\theta}$, markup θ
- set prices in Calvo contracts with avg. duration 4 quarters
- identical production functions $y_t = A_t \bar{k}^{1-\alpha} I_t^{\alpha}$

- face Dixit-Stiglitz demand with elasticity $\frac{1+\theta}{\theta}$, markup θ
- set prices in Calvo contracts with avg. duration 4 quarters
- identical production functions $y_t = A_t \bar{k}^{1-\alpha} I_t^{\alpha}$
- have firm-specific capital stocks

- face Dixit-Stiglitz demand with elasticity $\frac{1+\theta}{\theta}$, markup θ
- set prices in Calvo contracts with avg. duration 4 quarters
- identical production functions $y_t = A_t \bar{k}^{1-\alpha} I_t^{\alpha}$
- have firm-specific capital stocks
- face aggregate technology $A_t = \rho_A A_{t-1} + \varepsilon_t^A$

- face Dixit-Stiglitz demand with elasticity $\frac{1+\theta}{\theta}$, markup θ
- set prices in Calvo contracts with avg. duration 4 quarters
- identical production functions $y_t = A_t \bar{k}^{1-\alpha} I_t^{\alpha}$
- have firm-specific capital stocks
- face aggregate technology $A_t = \rho_A A_{t-1} + \varepsilon_t^A$

Parameters
$$\theta = .2$$
, $\rho_{A} = .9$, $\sigma_{A}^{2} = .01^{2}$

Government:



Government:

• imposes lump-sum taxes G_t on households

Government:

- imposes lump-sum taxes G_t on households
- destroys the resources it collects

Government:

- imposes lump-sum taxes G_t on households
- destroys the resources it collects
- $\bullet \ G_t = \rho_G G_{t-1} + \varepsilon_t^G$

Government:

Background/Motivation

- imposes lump-sum taxes G_t on households
- destroys the resources it collects
- $G_t = \rho_G G_{t-1} + \varepsilon_t^G$

Parameters $\rho_G = .9$, $\sigma_G^2 = .004^2$

Government:

- imposes lump-sum taxes G_t on households
- destroys the resources it collects

$$\bullet \ G_t = \rho_G G_{t-1} + \varepsilon_t^G$$

Parameters
$$\rho_G = .9$$
, $\sigma_G^2 = .004^2$

Monetary Authority:

Government:

Background/Motivation

- imposes lump-sum taxes G_t on households
- destroys the resources it collects
- $\bullet \ G_t = \rho_G G_{t-1} + \varepsilon_t^G$

Parameters $\rho_G = .9$, $\sigma_G^2 = .004^2$

Monetary Authority:

$$i_t = \rho_i i_{t-1} + (1 - \rho_i) [i^* + g_y(y_t - y_{t-1}) + g_\pi \pi_t] + \varepsilon_t^i$$

Government:

Background/Motivation

- imposes lump-sum taxes G_t on households
- destroys the resources it collects
- $G_t = \rho_G G_{t-1} + \varepsilon_t^G$

Parameters $\rho_G = .9$, $\sigma_G^2 = .004^2$

Monetary Authority:

$$i_t = \rho_i i_{t-1} + (1 - \rho_i) [i^* + g_y(y_t - y_{t-1}) + g_\pi \pi_t] + \varepsilon_t^i$$

Parameters $\rho_i = .7$, $g_v = 0.5$, $g_{\pi} = 2$, $\sigma_i^2 = .004^2$



In DSGE framework, convenient to work with a default-free *consol*, a perpetuity that pays \$1 (nominal) every period

In DSGE framework, convenient to work with a default-free *consol*, a perpetuity that pays \$1 (nominal) every period

Price of the consol:

$$p_t^{(\infty)} = 1 + E_t m_{t+1} p_{t+1}^{(\infty)}$$

In DSGE framework, convenient to work with a default-free *consol*, a perpetuity that pays \$1 (nominal) every period

Price of the consol:

$$p_t^{(\infty)}=1+E_tm_{t+1}p_{t+1}^{(\infty)}$$

Risk-neutral consol price:

$$p_t^{(\infty)rn} = 1 + e^{-i_t} E_t p_{t+1}^{(\infty)rn}$$

In DSGE framework, convenient to work with a default-free *consol*, a perpetuity that pays \$1 (nominal) every period

Price of the consol:

$$p_t^{(\infty)} = 1 + E_t m_{t+1} p_{t+1}^{(\infty)}$$

Risk-neutral consol price:

$$p_t^{(\infty)rn} = 1 + e^{-i_t} E_t p_{t+1}^{(\infty)rn}$$

Term premium:

Conclusions

In DSGE framework, convenient to work with a default-free *consol*, a perpetuity that pays \$1 (nominal) every period

Price of the consol:

Background/Motivation

$$p_t^{(\infty)} = 1 + E_t m_{t+1} p_{t+1}^{(\infty)}$$

Risk-neutral consol price:

$$p_t^{(\infty)rn} = 1 + e^{-i_t} E_t p_{t+1}^{(\infty)rn}$$

Term premium:

$$\log\left(\frac{p_t^{(\infty)}}{p_t^{(\infty)}-1}\right) - \log\left(\frac{p_t^{(\infty)rn}}{p_t^{(\infty)rn}-1}\right)$$

Solving the Model

The benchmark model above has a relatively large numer of state variables:

The benchmark model above has a relatively large numer of state variables: C_{t-1} , A_{t-1} , G_{t-1} , i_{t-1} , Δ_{t-1}

The benchmark model above has a relatively large numer of state variables: C_{t-1} , A_{t-1} , G_{t-1} , i_{t-1} , Δ_{t-1} , ε_t^A , ε_t^G , ε_t^i

Conclusions

The benchmark model above has a relatively large numer of state variables: C_{t-1} , A_{t-1} , G_{t-1} , i_{t-1} , Δ_{t-1} , ε_t^A , ε_t^G , ε_t^i

Value function iteration strategies are intractable

The benchmark model above has a relatively large numer of state variables: C_{t-1} , A_{t-1} , G_{t-1} , i_{t-1} , Δ_{t-1} , ε_t^A , ε_t^G , ε_t^i

Value function iteration strategies are intractable

The benchmark model above has a relatively large numer of state variables: C_{t-1} , A_{t-1} , G_{t-1} , i_{t-1} , Δ_{t-1} , ε_t^A , ε_t^G , ε_t^i

Value function iteration strategies are intractable

We solve the model by approximation around the nonstochastic steady state (perturbation methods)

In a first-order approximation, term premium is zero

The benchmark model above has a relatively large numer of state variables: C_{t-1} , A_{t-1} , G_{t-1} , i_{t-1} , Δ_{t-1} , ε_t^A , ε_t^G , ε_t^i

Value function iteration strategies are intractable

- In a first-order approximation, term premium is zero
- In a second-order approximation, term premium is a constant (sum of variances)

The benchmark model above has a relatively large numer of state variables: C_{t-1} , A_{t-1} , G_{t-1} , i_{t-1} , Δ_{t-1} , ε_t^A , ε_t^G , ε_t^G

Value function iteration strategies are intractable

- In a first-order approximation, term premium is zero
- In a second-order approximation, term premium is a constant (sum of variances)
- So we compute a third-order approximation of the solution around nonstochastic steady state

The benchmark model above has a relatively large numer of state variables: C_{t-1} , A_{t-1} , G_{t-1} , i_{t-1} , Δ_{t-1} , ε_t^A , ε_t^G , ε_t^i

Value function iteration strategies are intractable

- In a first-order approximation, term premium is zero
- In a second-order approximation, term premium is a constant (sum of variances)
- So we compute a third-order approximation of the solution around nonstochastic steady state
- perturbationAIM algorithm in Swanson, Anderson, Levin (2006) quickly computes nth order approximations



Impulse Responses

Figure 1

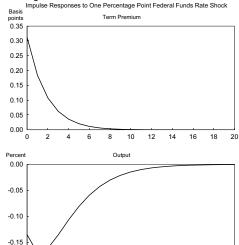
-0.20

2

6

10 12 14 16

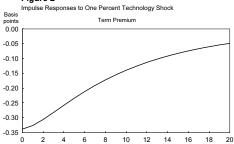
Quarters

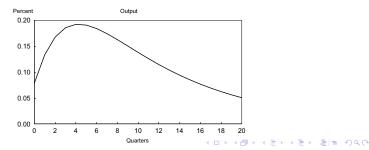




Impulse Responses

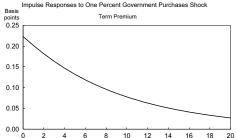
Figure 2

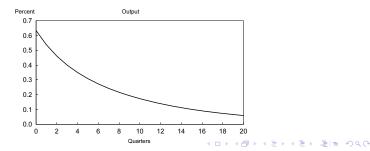




Impulse Responses

Figure 3





Theoretical Limitations:

 No consensus on how to model equity premium, risk premia in general

- No consensus on how to model equity premium, risk premia in general
- Term premia in benchmark New Keynesian model were very small, very stable

- No consensus on how to model equity premium, risk premia in general
- Term premia in benchmark New Keynesian model were very small, very stable
- Representative household assumption may work poorly for asset pricing

- No consensus on how to model equity premium, risk premia in general
- Term premia in benchmark New Keynesian model were very small, very stable
- Representative household assumption may work poorly for asset pricing
- Stochastic pricing kernel may not match standard utility functions used in macroeconomic models

Computational Limitations:

• Closed-form solutions exist only for simplest possible models

- Closed-form solutions exist only for simplest possible models
- Linearization or second-order approximation around nonstochastic steady state is not an option

- Closed-form solutions exist only for simplest possible models
- Linearization or second-order approximation around nonstochastic steady state is not an option
- Value function iteration is tractable only for very small models

- Closed-form solutions exist only for simplest possible models
- Linearization or second-order approximation around nonstochastic steady state is not an option
- Value function iteration is tractable only for very small models
- Medium-size New Keynesian models are required to match impulse responses of macroeconomic variables (CEE, ACEL)

- Closed-form solutions exist only for simplest possible models
- Linearization or second-order approximation around nonstochastic steady state is not an option
- Value function iteration is tractable only for very small models
- Medium-size New Keynesian models are required to match impulse responses of macroeconomic variables (CEE, ACEL)
- Large-scale models (GEM, SIGMA) becoming standard for macroeconomic policy analysis

Macro-Finance Analysis

- Macro-Finance Analysis
 - VAR-based Macro-Finance Models
 - New Keynesian Macro-Finance Models

Literature follows Ang and Piazzesi (2003)

Literature follows Ang and Piazzesi (2003)

State variables X_t follow a VAR:

$$X_t = \mu + \Phi X_{t-1} + \Sigma \varepsilon_t$$

Conclusions

Literature follows Ang and Piazzesi (2003)

State variables X_t follow a VAR:

$$X_t = \mu + \Phi X_{t-1} + \Sigma \varepsilon_t$$

Ad hoc stochastic pricing kernel:

$$m_{t+1} = \exp\left(-i_t - \frac{1}{2}\lambda_t'\lambda_t - \lambda_t'\varepsilon_{t+1}\right)$$

with

$$\lambda_t = \lambda_0 + \lambda_1 X_t$$

and ε_{t+1} conditionally log-normal

Appealing framework that allows changes in macroeconomic variables to affect term premium

Appealing framework that allows changes in macroeconomic variables to affect term premium

But ignores effect of term premium on macroeconomy:

Appealing framework that allows changes in macroeconomic variables to affect term premium

But ignores effect of term premium on macroeconomy:

 To maintain tractability, literature sharply restricts interaction between term premium and economic variables

Appealing framework that allows changes in macroeconomic variables to affect term premium

But ignores effect of term premium on macroeconomy:

- To maintain tractability, literature sharply restricts interaction between term premium and economic variables
- In Ang-Piazzesi and Bernanke-Reinhart-Sack (2005), term premium assumed to have no effect on economy

Appealing framework that allows changes in macroeconomic variables to affect term premium

But ignores effect of term premium on macroeconomy:

- To maintain tractability, literature sharply restricts interaction between term premium and economic variables
- In Ang-Piazzesi and Bernanke-Reinhart-Sack (2005), term premium assumed to have no effect on economy
- In Ang-Piazzesi-Wei (2006), term premium assumed to have same effect on economy as changes in risk-neutral rate

State variables X_t follow a linearized New Keynesian system of structural equations instead of a VAR

State variables X_t follow a linearized New Keynesian system of structural equations instead of a VAR

Gives the model more structure while retaining tractability, empirical fit

State variables X_t follow a linearized New Keynesian system of structural equations instead of a VAR

Gives the model more structure while retaining tractability, empirical fit

As with VAR-based models, though, effects of term premium on economy are assumed rather than estimated or derived:

State variables X_t follow a linearized New Keynesian system of structural equations instead of a VAR

Gives the model more structure while retaining tractability, empirical fit

As with VAR-based models, though, effects of term premium on economy are assumed rather than estimated or derived:

Linearized IS curve allows no role for term premium

State variables X_t follow a linearized New Keynesian system of structural equations instead of a VAR

Gives the model more structure while retaining tractability, empirical fit

As with VAR-based models, though, effects of term premium on economy are assumed rather than estimated or derived:

- Linearized IS curve allows no role for term premium
- Rudebusch-Wu (2004) allow for latent factors to affect economy, but in effect assume that effect of term premium and risk-neutral rate are the same

Reduced-Form Analysis

- Reduced-Form Analysis
 - The Yield Curve Slope and Forecasting GDP
 - Five Measures of the Term Premium
 - Importance of Term Premium for Forecasting GDP

A large literature uses slope of yield curve to forecast GDP:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2(i_t^{(n)} - i_t) + \varepsilon_t$$

A large literature uses slope of yield curve to forecast GDP:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2(i_t^{(n)} - i_t) + \varepsilon_t$$

Note: This is a reduced-form forecasting equation, no structure

A large literature uses slope of yield curve to forecast GDP:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2(i_t^{(n)} - i_t) + \varepsilon_t$$

Note: This is a reduced-form forecasting equation, no structure

Motivation: $i_t^{(n)}$ proxies for i^* , so $i_t^{(n)} - i_t$ proxies for stance of monetary policy

A large literature uses slope of yield curve to forecast GDP:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2(i_t^{(n)} - i_t) + \varepsilon_t$$

Note: This is a reduced-form forecasting equation, no structure

Motivation: $i_t^{(n)}$ proxies for i^* , so $i_t^{(n)} - i_t$ proxies for stance of monetary policy

Estimates in literature consistently find $\beta_2 > 0$, highly significant

If $i_t^{(n)}$ proxies for i^* , then:

If $i_t^{(n)}$ proxies for i^* , then:

ullet expectations component of $i_t^{(n)}$ should be better measure of i^*

If $i_t^{(n)}$ proxies for i^* , then:

- ullet expectations component of $i_t^{(n)}$ should be better measure of i^*
- term premium itself might have predictive power for GDP

If $i_t^{(n)}$ proxies for i^* , then:

- ullet expectations component of $i_t^{(n)}$ should be better measure of i^*
- term premium itself might have predictive power for GDP

Separate yield curve slope $i_t^{(n)} - i_t$ into:

If $i_t^{(n)}$ proxies for i^* , then:

- expectations component of $i_t^{(n)}$ should be better measure of i^*
- term premium itself might have predictive power for GDP

Separate yield curve slope $i_t^{(n)} - i_t$ into:

$$\underbrace{\left(\frac{1}{n}\sum_{j=0}^{n-1}E_{t}i_{t+j}-i_{t}\right)}_{\text{expectations component}} + \underbrace{\left(i_{t}^{(n)}-\frac{1}{n}\sum_{j=0}^{n-1}E_{t}i_{t+j}\right)}_{\text{term premium}}$$

If $i_t^{(n)}$ proxies for i^* , then:

Background/Motivation

- expectations component of $i_t^{(n)}$ should be better measure of i^*
- term premium itself might have predictive power for GDP

Separate yield curve slope $i_t^{(n)} - i_t$ into:

$$\underbrace{\left(\frac{1}{n}\sum_{j=0}^{n-1}E_ti_{t+j}-i_t\right)}_{\textbf{exsp}_t} + \underbrace{\left(i_t^{(n)}-\frac{1}{n}\sum_{j=0}^{n-1}E_ti_{t+j}\right)}_{\textbf{tp}_t}$$

Generalize basic GDP forecasting equation to:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2 \exp_t + \beta_3 tp_t + \varepsilon_t$$

Generalize basic GDP forecasting equation to:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2 \exp_t + \beta_3 tp_t + \varepsilon_t$$

Forecasts using only yield curve slope in effect impose $\beta_2=\beta_3$

Generalize basic GDP forecasting equation to:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2 \exp_t + \beta_3 tp_t + \varepsilon_t$$

Forecasts using only yield curve slope in effect impose $\beta_2=\beta_3$

First paper to separate out term premium and investigate importance for forecasting is Hamilton-Kim (2002)

Generalize basic GDP forecasting equation to:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2 \exp_t + \beta_3 tp_t + \varepsilon_t$$

Forecasts using only yield curve slope in effect impose $\beta_2=\beta_3$

First paper to separate out term premium and investigate importance for forecasting is Hamilton-Kim (2002)

Generally, authors find:

Generalize basic GDP forecasting equation to:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2 \exp_t + \beta_3 tp_t + \varepsilon_t$$

Forecasts using only yield curve slope in effect impose $\beta_2=\beta_3$

First paper to separate out term premium and investigate importance for forecasting is Hamilton-Kim (2002)

Generally, authors find:

• $\beta_2 > 0$, highly significant

Generalize basic GDP forecasting equation to:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2 \exp_t + \beta_3 tp_t + \varepsilon_t$$

Forecasts using only yield curve slope in effect impose $\beta_2 = \beta_3$

First paper to separate out term premium and investigate importance for forecasting is Hamilton-Kim (2002)

Generally, authors find:

- $\beta_2 > 0$, highly significant
- $\beta_2 > \beta_3$ (can reject $\beta_2 = \beta_3$)

Generalize basic GDP forecasting equation to:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2 \exp_t + \beta_3 tp_t + \varepsilon_t$$

Forecasts using only yield curve slope in effect impose $\beta_2=\beta_3$

First paper to separate out term premium and investigate importance for forecasting is Hamilton-Kim (2002)

Generally, authors find:

- $\beta_2 > 0$, highly significant
- $\beta_2 > \beta_3$ (can reject $\beta_2 = \beta_3$)
- $\beta_3 > 0$, insignificant

- VAR
 - use VAR to construct risk-neutral 10-year yield
 - term premium is residual

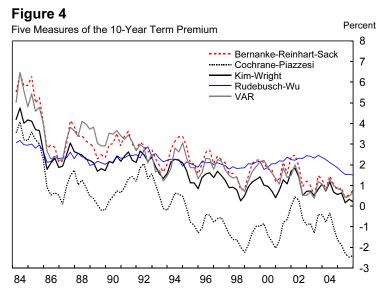
- VAR
 - use VAR to construct risk-neutral 10-year yield
 - term premium is residual
- Bernanke-Reinhart-Sack (2005)
 - VAR imposing no-arbitrage restrictions

- VAR
 - use VAR to construct risk-neutral 10-year yield
 - term premium is residual
- Bernanke-Reinhart-Sack (2005)
 - VAR imposing no-arbitrage restrictions
- Rudebusch-Wu (2004)
 - New Keynesian model imposing no-arbitrage

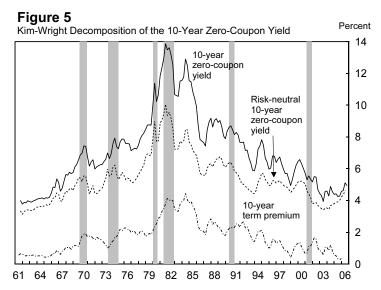
- VAR
 - use VAR to construct risk-neutral 10-year yield
 - term premium is residual
- Bernanke-Reinhart-Sack (2005)
 - VAR imposing no-arbitrage restrictions
- Rudebusch-Wu (2004)
 - New Keynesian model imposing no-arbitrage
- 4 Kim-Wright (2006)
 - no-arbitrage three-factor finance model

- VAR
 - use VAR to construct risk-neutral 10-year yield
 - term premium is residual
- Bernanke-Reinhart-Sack (2005)
 - VAR imposing no-arbitrage restrictions
- Rudebusch-Wu (2004)
 - New Keynesian model imposing no-arbitrage
- Kim-Wright (2006)
 - no-arbitrage three-factor finance model
- Cochrane-Piazzesi (2005)
 - excess return forecasting factor





Kim-Wright Term Premium



GDP Forecasting Results

Table 2 Prediction Equations for GDP Growth dependent variable: $y_{t+4} - y_t$

	1962-2005 Sample			
	(1)	(2)		
$y_t - y_{t-4}$	0.15 (1.57)	0.12 (1.18)		
$i_t^{(n)} - i_t$	0.64 (3.64)			
exsp _t		0.68 (4.03)		
tp _t		0.30 (0.92)		

GDP Forecasting Results

Table 2
Prediction Equations for GDP Growth dependent variable: $y_{t+4} - y_t$

	1962–2005 Sample			
	(1)	(2)		
$y_t - y_{t-4}$	0.15 (1.57)	0.12 (1.18)		
$i_t^{(n)}-i_t$	0.64 (3.64)			
exsp _t		0.68 (4.03)		
tp _t		0.30 (0.92)		

Note: we cannot reject hypothesis that coefficients on $exsp_t$, tp_t are equal

Regression Specification

Recall new Keynesian IS curve:

$$y_t = -\frac{1}{\gamma} E_t \sum_{j=0}^{\infty} \beta^j (i_{t+j} - \pi_{t+1+j}) + \varepsilon_t$$

Regression Specification

Recall new Keynesian IS curve:

$$y_t = -\frac{1}{\gamma} E_t \sum_{j=0}^{\infty} \beta^j (i_{t+j} - \pi_{t+1+j}) + \varepsilon_t$$

Suggests that yield spread should be related to *level* of GDP, rather than growth rate

Regression Specification

Recall new Keynesian IS curve:

$$y_t = -\frac{1}{\gamma} E_t \sum_{j=0}^{\infty} \beta^j (i_{t+j} - \pi_{t+1+j}) + \varepsilon_t$$

Suggests that yield spread should be related to *level* of GDP, rather than growth rate

To account for nonstationarity, forecasting regression specification should then be:

$$(y_{t+4} - y_t) = \beta_0 + \beta_1(y_t - y_{t-4}) + \beta_2(exsp_t - exsp_{t-4}) + \beta_3(tp_t - tp_{t-4}) + \varepsilon_t$$

GDP Forecasting Results

Table 2 (cont.)
Prediction Equations for GDP Growth dependent variable: $y_{t+4} - y_t$

	1962-2005 Sample			
	(3)	(4)		
$y_t - y_{t-4}$	0.32 (3.04)	0.38 (4.22)		
exsp _t	1.03 (5.64)			
$exsp_{t-4}$	-0.79 (-3.49)			
tp_t	-0.61 (-1.34)			
$ tp_{t-4} $	0.54 (1.24)			
$exsp_t - exsp_{t-4}$		0.96 (5.62)		
$tp_t - tp_{t-4}$		-0.77 (-1.95)		

Background/Motivation

	1962–2005 Sample			
	(3)	(4)		
$y_t - y_{t-4}$	0.32 (3.04)	0.38 (4.22)		
exsp _t	1.03 (5.64)			
$exsp_{t-4}$	-0.79 (-3.49)			
tp_t	-0.61 (-1.34)			
$ tp_{t-4} $	0.54 (1.24)			
$exsp_t - exsp_{t-4}$		0.96 (5.62)		
$tp_t - tp_{t-4}$		-0.77 (-1.95)		

Note: we strongly reject hypothesis that coefficients on $exsp_t$, tp_t are equal

- There is no structural, causal relationship running from the term premium to the economy
 - correlation is different for different structural shocks

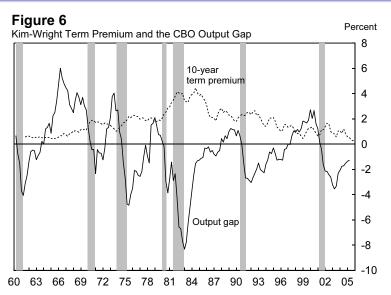
- There is no structural, causal relationship running from the term premium to the economy
 - correlation is different for different structural shocks
- Reduced-form evidence strongly suggests that policymakers should take term premium into account when forecasting
 - strongly rejected hypothesis that $\beta_3 = \beta_2$ in forecasting regression

- There is no structural, causal relationship running from the term premium to the economy
 - correlation is different for different structural shocks
- Reduced-form evidence strongly suggests that policymakers should take term premium into account when forecasting
 - strongly rejected hypothesis that $\beta_3=\beta_2$ in forecasting regression
- Oeclines in the term premium have typically been followed by economic expansion
 - true in both the post-1960 and post-1985 periods

Policymakers were right to closely watch declining term premium in 2004-5

- Policymakers were right to closely watch declining term premium in 2004-5
- Some reduced-form evidence that the Practitioner/Chairman View of macroeconomic implications of declining term premium was correct

Kim-Wright Term Premium and Output Gap



Cochrane-Piazzesi Term Premium Measure

Figure 1
Term Premium for Ten-Year Treasury Security
Implied by Cochrane-Piazzesi Results



Cochrane-Piazzesi Term Premium Measure

Figure 2
Comparison of Term Premium and One-Year Expected Excess Returns for Ten-Year Treasury Security

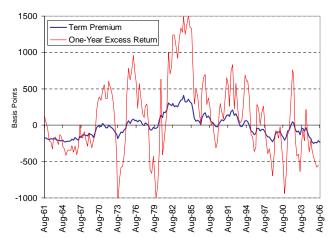


Table 1
Correlations between Five Measures of the Term Premium

	BRS	RW	KW	CP	VAR
BRS	1.00				
RW	0.76	1.00			
KW	0.98	0.81	1.00		
CP	0.92	0.87	0.96	1.00	
VAR	0.96	0.68	0.94	0.88	1.00